

# **MATH-505A: Homework # 4**

Due on Friday, September 19, 2014

**Saket Choudhary**  
**2170058637**

## Contents

<b>Exercise # 2.1</b>	<b>3</b>
(1) . . . . .	3
(2) . . . . .	3
(3) . . . . .	3
(4) . . . . .	4
(5) . . . . .	5
<b>Exercise # 2.3</b>	<b>6</b>
(1) . . . . .	6
(2) . . . . .	6
(3) . . . . .	6
(4) . . . . .	6
(5) . . . . .	7

**Exercise # 2.1****(1)****Given:**  $X$  is a random variable  $\implies$ 

$$\{\omega \in \Omega : X(\omega) \leq x\} \in \mathcal{F} \quad \forall x \in R \quad (1)$$

**Part A)** To Prove:  $aX$  is a random variableConsider  $Y = aX$ , then since equation 1 holds:**Case1:**  $a \geq 0$ Then  $\{\omega \in \Omega : aX(\omega) \leq x'\} \in \mathcal{F} \quad \forall x' \in R$  where  $x' = ax$ **Case2:**  $a \leq 0$ Then  $\{\omega \in \Omega : aX(\omega) \geq x'\} \forall x' \in R$  where  $x' = ax \implies \cup\{\{\omega \in \Omega : aX(\omega) \leq x''\}\}^c \in \mathcal{F}$  where  $x'' = x'$ **Case3:**  $a$  is 0Then,  $aX = 0$ **Case i:**  $x < 0$ 

$$\{\omega \in \Omega : aX(\omega) = \phi\} \in \mathcal{F}$$

**Case ii:**  $x \geq 0$ 

$$\{\omega \in \Omega : aX(\omega) = \Omega\} \in \mathcal{F}$$

Thus from all the above cases.

**Part (b):**Consider  $Y = X - X$ , Then:

$$Y = X(\omega) - X(\omega) \forall \omega \in R \implies Y = 0$$

Consider  $Y = X + X$ , Then  $Y = X(\omega) + X(\omega) \forall \omega \in \Omega \implies Y = 2X(\omega) \forall \omega \in \Omega$  Thus  $Y = 2X$ .**(2)**For part 1,  $Y' = aX$  is also a random variable:**To Find:** Distribution function of  $Y = aX + b$ 

$$\text{Consider } P(Y \leq y) = P(aX + b \leq y) = P(X \leq \frac{y-b}{a}) \implies P(Y \leq y) = P(X \leq \frac{y-b}{a})$$

**(3)**

$$p(H) = p; p(T) = 1 - p$$

Tossing a coin  $n$  times is a binomial process (each individual toss is a bernoulli process) and let A be the event such that  $k$  out of  $n$  tosses are heads and this can occur in  $\binom{n}{k}$  ways with probability  $p^k$ . There would also be  $n - k$  tails and the probability for that is  $(1 - p)^{n-k}$ . Thus,:

$$p(A) = \binom{n}{k} p^k * (1 - p)^{n-k}$$

$$\text{For a fair coin, } p = \frac{1}{2} \text{ and hence } p(A) = \binom{n}{k} \left(\frac{1}{2}\right)^k \left(\frac{1}{2}\right)^{n-k} = \binom{n}{k} \left(\frac{1}{2}\right)^n$$

(4)

A distribution function satisfies the following set of properties:

a)  $\lim_{x \rightarrow -\infty} F(x) = 0, \lim_{x \rightarrow \infty} F(x) = 1$

b) if  $x < y$  then  $F(x) \leq F(y)$ ,

c) F is right continuous,  $c < x < c + \delta$  then  $|F(x) - F(c)| < \epsilon$  for  $\epsilon > 0, \delta > 0$

Consider  $Y = \lambda F + (1 - \lambda)G$ , Both G,F satisfy a, b, c Then  $\lim_{x \rightarrow -\infty} Y(x) = \lambda \lim_{x \rightarrow -\infty} F(x) + (1 - \lambda) \lim_{x \rightarrow -\infty} G(x) \implies \lim_{x \rightarrow -\infty} Y(x) = 0$

Similarly considering limit as  $x \rightarrow \infty$ : Then  $\lim_{x \rightarrow \infty} Y(x) = \lambda \lim_{x \rightarrow \infty} F(x) + (1 - \lambda) \lim_{x \rightarrow \infty} G(x) \implies \lim_{x \rightarrow \infty} Y(x) = \lambda * 1 + (1 - \lambda) * 1 = 1$

Since for  $x < y$ , then  $F(x) < F(y); G(x) < G(y) \implies \lambda F(x) < \lambda F(y); (1 - \lambda)G(x) < (1 - \lambda)G(y)$  since  $0 \leq \lambda \leq 1$

Adding the two inequalities we get:

$$\lambda F(x) + (1 - \lambda)G(x) < \lambda F(y) + (1 - \lambda)G(y) \implies Y(x) < Y(y).$$

Since F,G are right continuous, any linear combination of these would be right continuous too.

Hence  $Y = \lambda F + (1 - \lambda)G$  satisfies all the 3 required properties and is a distribution function.

(5)

Since  $F$  is a distribution function:

(i)  $\lim_{x \rightarrow -\infty} F(x) = 0; \lim_{x \rightarrow \infty} F(x) = 1$

(ii) If  $x < y$  then,  $F(x) < F(y)$

(iii)  $F$  is right continuous

**Part a)**  $F(x)^r$

(i)  $\lim_{x \rightarrow -\infty} F(x)^r = 0$  since  $\lim_{x \rightarrow -\infty} F(x) = 0$  and  $r > 0$

(ii) If  $x < y$  as  $F(x) < F(y)$  and  $r > 0 \implies F(x)^r < F(y)^r$

(iii) Since  $r > 0$  and  $F(x)$  is right-continuous  $F(x)^r$  is right continuous. (One possible case where  $F(x)^r$  would not have been right continuous is for  $r < 0$  say  $r = -1$  where  $F(x)^{-1}$  is not right continuous at all  $x_0$  such that  $F(x_0) = 0$ .)

**Part b)**  $1 - (1 - F(x))^r$

(i)  $\lim_{x \rightarrow -\infty} (1 - (1 - F(x))^r) = 1 - \lim_{x \rightarrow -\infty} (1 - F(x))^r = 1 - (1 - 0)^r = 0$

Similarly for  $\lim_{x \rightarrow \infty} (1 - (1 - F(x))^r) = 1 - (1 - 1)^r = 1$

(ii) If  $x < y$ ,  $F(x) < F(y) \implies -F(x) > -F(y) \implies 1 - F(x) > 1 - F(y) \implies (1 - F(x))^r > (1 - F(y))^r \forall r > 0$  Thus,  $1 - (1 - F(x))^r < 1 - (1 - F(y))^r$

(iii) Since  $F(x)$  is right continuous,  $1 - F(x)$  is right continuous  $\implies (1 - F(x))^r$  is right continuous (since  $r > 0$ ) implies  $1 - (1 - F(x))^r$  is right continuous

**Part c)**  $F(x) + (1 - F(x))\log(1 - F(x))$

(i)  $\lim_{x \rightarrow -\infty} (F(x) + (1 - F(x))\log(1 - F(x))) = \lim_{x \rightarrow -\infty} F(x) + \lim_{x \rightarrow -\infty} (1 - F(x))\log(1 - F(x)) = 0 + (1 - 0)\log(1 - 0) = 0$

Consider  $\lim_{x \rightarrow \infty} (F(x) + (1 - F(x))\log(1 - F(x))) = \lim_{x \rightarrow \infty} F(x) + \lim_{x \rightarrow \infty} (1 - F(x))\log(1 - F(x)) = 1 + (1 - 1)\log(1 - 1) = 0$

(ii) If  $x < y$  then  $F(x) < F(y) \implies 1 - F(x) > 1 - F(y)$ . Since  $\log$  is a monotonic non-increasing function in  $[0, 1]$  and is in fact negative definite:  $\log(1 - F(x)) < \log(1 - F(y))$  and  $1 - F(x) > 1 - F(y) \implies -F(x)\log(1 - F(x)) < -F(y)\log(1 - F(y))$  (This holds only because  $\log$  is negative definite in  $[0, 1]$ ) Thus,  $F(x) - F(x)\log(1 - F(x)) < F(y) - F(y)\log(1 - F(y))$

(iii)  $F(x) - F(x)\log(1 - F(x))$  is right continuous as  $\log(1 - F(x))$  is right continuous

**Part d)**  $(F(x) - 1)e + \exp(1 - F(x))$

(i)  $\lim_{x \rightarrow -\infty} (F(x) - 1)e + \exp(1 - F(x)) = \lim_{x \rightarrow -\infty} (F(x) - 1)e + \exp(\lim_{x \rightarrow -\infty} (1 - F(x))) = (0 - 1)e + \exp(1 - 0) = -e + e = 0$

$\lim_{x \rightarrow \infty} (F(x) - 1)e + \exp(1 - F(x)) = \lim_{x \rightarrow \infty} (F(x) - 1)e + \exp(\lim_{x \rightarrow \infty} (1 - F(x))) = (1 - 1)e + \exp(1 - 1) = 0 + 1 = 1$

(ii) if  $x < y$ ,  $F(x) < F(y) \implies F(x) - 1 < F(y) - 1 \implies (F(x) - 1)e < (F(y) - 1)e$  Also,  $1 - F(x) > 1 - F(y)$

Since  $\exp$  is a non-increasing function in  $[0, 1]$   $\exp(1 - F(x)) < \exp(1 - F(y))$

Thus,

$(F(x) - 1)e + \exp(1 - F(x)) < (F(y) - 1)e + \exp(1 - F(y))$

(iii)  $(F(x) - 1)e + \exp(1 - F(x))$  is right continuous as  $\exp$  is right continuous.

$FG$  is also a density function since it satisfies:

(i)  $\lim_{x \rightarrow -\infty} F(x)G(x) = \lim_{x \rightarrow -\infty} F(x) * \lim_{x \rightarrow -\infty} G(x) = 0$

And  $\lim_{x \rightarrow \infty} F(x)G(x) = \lim_{x \rightarrow \infty} F(x) * \lim_{x \rightarrow \infty} G(x) = 1$

(ii) If  $x < y$ ,  $F(x) < F(y)$  and  $G(x) < G(y) \implies F(x)G(x) < F(y)G(y)$

(iii) Since  $F(x), G(x)$  are right continuous

## Exercise # 2.3

(1)

**Given:**  $\lim_{m \rightarrow -\infty} a_m \rightarrow -\infty$  and  $\lim_{m \rightarrow \infty} a_m \rightarrow \infty$ ;  $G(x) = P(X \leq a_m)$  when  $a_{m-1} \leq x \leq a_m$ ;  $a_m$  is a strictly increasing sequence.

Sequence  $a$  is chosen so that  $\sup_m |a_m - a_{m-1}|$  becomes smaller and smaller so even though the sequence is increasing the successive difference between the terms keep on decreasing essentially indicating  $a_m$  saturates as  $m \rightarrow \infty$

so  $F(x)$  approaches  $G(x) \forall x$

(2)

**Given:**  $g(x)$  is continuous and strictly increasing,  $X$  is a random variable:

Since  $g(X)$  is continuous and strictly increasing  $\implies g^{-1}$  exists.

Consider  $\{Y \leq y\} \implies \{g(X) \leq y\}$ . Since  $g^{-1}$  exists, such a set is equivalent to:  $\{X \leq g^{-1}(y)\}$  which belongs to  $\mathcal{F}$  as  $g: R \rightarrow R$ . Since the last equations hold true and hence  $\{Y \leq y\} \in \mathcal{F}$  so  $g(X)$  is a RV

(3)

$$F(x) = P(X \leq x) = \begin{cases} 0 & \text{if } x \leq 0, \\ x & \text{if } 0 < x \leq 1, \\ 1 & \text{if } x > 1, \end{cases}$$

**To Prove:**  $Y = F^{-1}(x)$  is a random variable: Consider  $\{Y \leq$

$y\} = \{F^{-1}(x) \leq y\}$ . Since  $F$  is continuous and strictly increasing  $\implies F^{-1}(x)$  exists in  $R$  so  $\{Y \leq y\} = \{F^{-1}(x) \leq y\} = \{x \leq F(y)\} = \{x \leq P(X \leq y)\} \in \mathcal{F}$

Since the last equation holds true, it  $\implies \{Y \leq y\} \in \mathcal{F} \implies Y = F^{-1}$  is a RV.

$F$  should necessarily be continuous and monotonic for the inverse to exist!

(4)

$f, g$  are density functions:

$$\int_{-\infty}^{\infty} f(x) dx = 1 \text{ and } \int_{-\infty}^{\infty} g(x) dx = 1$$

Consider  $y(x) = \lambda f(x) + (1 - \lambda)g(x)$  Thus,  $\int_{-\infty}^{\infty} y(x) dx = \int_{-\infty}^{\infty} \lambda f(x) dx + \int_{-\infty}^{\infty} (1 - \lambda)g(x) dx = \lambda * \int_{-\infty}^{\infty} f(x) dx + (1 - \lambda) * \int_{-\infty}^{\infty} g(x) dx = \lambda * 1 + (1 - \lambda) * 1 = 1$

Thus  $\lambda f(x) + (1 - \lambda)g(x)$  is a density function too.

Now consider  $y(x) = f(x)g(x)$ , then :

$$\int_{-\infty}^{\infty} y(x) dx = \int_{-\infty}^{\infty} f(x)g(x) dx \text{ Clearly this is not necessarily equal to 1 so } fg \text{ is not a density function!}$$

(5)

$$\mathbf{Part a)} f(x) = \begin{cases} cx^{-d} & x > 1, \\ 0 & \text{otherwise,} \end{cases}$$

$F(x) = \int_{-\infty}^{\infty} f(c)dx = \int_1^{\infty} cx^{-d} = 1 \implies \frac{-c}{-d+1} = 1 \implies c = d - 1$  and  $-d + 1 \leq 1$  i.e  $d \geq 0$  else the integral blows up to  $\infty$

Thus for  $d \geq 0$  this is a density function satisfying  $c = d - 1$

$$\mathbf{Part b)} f(x) = ce^x(1 + e^x)^{-2} x \in R$$

$F(x) = \int_{-\infty}^{\text{infy}} ce^x(1 + e^x)^{-2}$  Let  $t = e^x + 1$  then  $e^x dx = dt$   $F(x) = \int_1^{\infty} ct^{-2} dt$   $F(x = 1) = -(c * 0 - c) = 1$

Thus  $c = 1$ .