# CSCI-567: Assignment  $\#$  2

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# Contents



# Todo list

# Problem 1

Problem 1: (a)

Linear regression assumes that the regressors have been observed 'truly' and as such the dependent variables Y are the ones that are uncertain. The analogy is simpler to think when Y is a 'response', caused due to some indepndent variable X. hence though X is measured absolutely, (dependent varible) Y's measurement should be accounted for errors.

## Problem 1: (b)

In order to make linear regression robust to outliers, a näive solution will choose "absolute deviation" $(L1)$ norm) over "squared error" $(L2$  norm) as the criterion for loss function. The reason this might work out in most cases(especially when the outliers belong to a non normal distribution) is that "squared error" will blow up errors when they are large. Thus L2 norm will give more weight to large residuals( $|y - w^T x|^2$ )  $|y - w^T x|$  and we are trying to minimise this error), while the L1 norm gives equal weights to all residuals.

#### Problem 1: (c)

A quick way to realise this is to consider the "scale" of any two independent variables. Say one of the dependent variables is 'time'. Rescaling time from hours to seconds will also rescale its coefficient(approximately by a factor of 60), but the importance remains the same!

Another example is to consider a model with two dependent variables that affect the dependent variable in a similar manner (or are equally important regressors). However they are on different scales say  $X1$  on  $[1-100]$  and X2 on  $[0-1]$ , resulting in the coefficient of X1 being too smaller than that of X2 in a linear regression setting.

#### Problem 1: (d)

If the dependent variables are perfect linear combination, the matrix  $XX<sup>T</sup>$  will be non invertible.

#### Problem 1: (e)

A simple solution would be to use  $k - 1$  bits instead of k bits for k categories. For example, using the following setup for a 3 category setup:

> $Red = 0$  $Green = 10$  $Blue = 0.1$

Here is an alternate solution that exploits the property of features still being equidistant from the origin(though they are no longer equidistant from each other)

### Problem 1: (f)

If the independent variables are highly correlated, the coefficients might still be entirely different. From the example in Part (c) abov

#### Problem 1: (g)

Using a posterior probability cutoff of 0.5 in linear regression is not same as 0.5 for logistic. A 0.5 rehsold on logistic guarantees that the point all points lying to the right belong to one class. However for a regression problem, this is not true, because the predicted value of y is an 'intepolated or extraplolated' In any case, logistic regression is a better choice, since the output is constrained in the range of  $0 - 1$ , which can be treated directly as a probability values as compared to the less intuitive relation with the output of the linear regression.

### Problem 1: (h)

When the number of variables exceed the number of samples, the system is undetermined. And yes, it can be solved by simply obtaining psuedo-inverse of  $X$  which is always defined.

# Problem 2

### Problem 2: (a)

Class 1:  $\vec{x} = (x_1, x_2, \cdots x_{2D})$  where each  $x_i \sim N(0, \sigma^2)$ Class 2:  $\vec{x} = (x_1, x_2, \cdots x_D, x_{D+1} + \delta, \cdots x_{2D})$ From first principles, the discriminant curve is given by:

$$
P(y=1|x) \ge P(y=0|x)
$$

And hence we have:

$$
\log(P(y=1|x)) \ge \log(P(y=0|x))
$$
  
\n
$$
\log(P(x|y=1)p(y=1) \ge \log(P(x|y=0)p(y=0))
$$
  
\n
$$
\log(P(x|y=1)) + \log(P(y=1)) \ge \log(P(x|y=0)) + \log(p(y=0))
$$
\n(1)

Now, since  $x$  is 2D dimensional and assuming independece of all attributes:

$$
P(x|y=1) = \prod_{i=1}^{2D} p(x_i|y=1)
$$
  

$$
\log(P(x|y=1)) = \sum_{i=1}^{2D} \log(p(x_i|y=1))
$$
  

$$
\log(P(x|y=1)) = -D \log(2\pi\sigma^2) - \sum_{i=1}^{2D} \frac{x_i^2}{2\sigma^2}
$$
 (2)

Similarly for class 0:  $x_i \sim N(0, \sigma^2) \forall x \in \{1..D\}$  and  $x_i \sim N(\delta, \sigma^2) \forall x \in \{D+1..2D\}$  Notice that the latter is a shifted normal.

$$
P(x|y=0) = \prod_{i=1}^{2D} p(x_i|y=0)
$$
  
\n
$$
\log(P(x|y=0) = \sum_{i=1}^{D} \log(p(x_i|y=1)) + \sum_{i=1}^{D} \log(p(x_i|y=1))
$$
  
\n
$$
\log(P(x|y=0)) = -D \log(2\pi\sigma^2) - \sum_{i=1}^{D} \frac{x_i^2}{2\sigma^2} - \sum_{i=D+1}^{2D} \frac{(x_i - \delta)^2}{2\sigma^2}
$$
\n(3)

Plugging  $(2), (3)$  in  $(1)$  we get:

$$
-D \log(2\pi\sigma^2) - \sum_{i=1}^{2D} \frac{x_i^2}{2\sigma^2} + \log(p(y=1)) \ge -D \log(2\pi\sigma^2) - \sum_{i=1}^{D} \frac{x_i^2}{2\sigma^2} - \sum_{i=D+1}^{2D} \frac{(x_i - \delta)^2}{2\sigma^2} + \log(p(y=0))
$$
  

$$
\log(p(y=1)) + \sum_{i=D+1}^{2D} \frac{x_i^2}{2\sigma^2} \ge -\sum_{i=D+1}^{2D} \frac{x_i^2 - 2\delta x_i + \delta^2}{2\sigma^2} + \log(p(y=0))
$$
  

$$
\log(p(y=1)) - \log(p(y=0)) \ge \frac{D\delta}{\sigma^2} \sum_{i=D+1}^{2D} x_i - D\frac{\delta^2}{2\sigma^2}
$$
  

$$
\log(p(y=1)) - \log(p(y=0)) \ge \frac{D\delta}{\sigma^2} \sum_{i=1}^{D} x_i - D\frac{\delta^2}{2\sigma^2}
$$
  

$$
-\frac{D\delta}{\sigma^2} \sum_{i=1}^{D} x_i + D\frac{\delta^2}{2\sigma^2} + \log(p(y=1)) - \log(p(y=0)) \ge 0
$$

Where the change of indices in the penultimate step is permitted since  $x_i$  are i.i.d(after taking care of the shifted mean)

Now consider the general form solution of LDA and GDA:

$$
\sum_{i} b_i x_i + c \ge 0
$$
 (LDA)

$$
\sum_{i} a_i x_i^2 + \sum b_i x_i + c \ge 0
$$
 (GDA)

Where  $x_i$  represents the  $i^{th}$  dimension indepdent variable, where each  $x_i$  is a feature/attribute and hence is not limited to 2 dimensional special case.

In this case, owing to the homoscedasticity assumption(the variance of the two class conditions being equal to  $\sigma^2$ ) LDA and GDA return the same solution. Solution form for  $LDA$ :

$$
-\frac{D\delta}{\sigma^2} \sum_{i=1}^{D} x_i + D \frac{\delta^2}{2\sigma^2} + \log(p(y=1)) - \log(p(y=0)) \ge 0
$$

Assuming equal priors,  $p(y = 1) = p(y = 0)$ ,

$$
-\frac{D\delta}{\sigma^2} \sum_{i=1}^{D} x_i + D \frac{\delta^2}{2\sigma^2} \ge 0
$$

$$
-\sum_{i=1}^{D} x_i + \frac{\delta}{2} \ge 0
$$

and hence for LDA  $b_i = 0 \forall i \in \{1..D\}$  and  $b_i = -\frac{D\delta}{\sigma^2} \forall i \in \{D+1..2D\}$  (simplifies to  $-1$  for the case with equal priors). and  $c = \frac{D\delta^2}{2\sigma^2}$  (simplifies to  $\frac{\delta}{2}$  for the case with equal priors) In either case it does depend on  $\delta$ Solution for  $GDA$ :

$$
-\frac{D\delta}{\sigma^2} \sum_{i=1}^{D} x_i + D \frac{\delta^2}{2\sigma^2} + \log(p(y=1)) - \log(p(y=0)) \ge 0
$$

and hence  $a_i = 0 \forall i$  and  $b_i = 0 \forall i \in \{1..D\}$  and  $b_i = -\frac{D\delta}{\sigma^2} \forall i \in \{D+1..2D\}$  (simplifies to  $-1$  for the case with equal priors).

In either case it does depend on  $\delta$ 

### Problem 2: (b1)

$$
P(X|Y = c_1) \sim N(\mu_1, \Sigma) \text{ and } p(X|Y = c_2) \sim N(\mu_2, \Sigma)
$$
  
\nwhere  $\mu_1, \mu_2 \in R^D, \Sigma \in R^{D \times D}$   
\n
$$
P(Y = 1|X) = \frac{P(X|Y = 1)P(Y = 1)}{P(X)}
$$
\n
$$
= \frac{P(X|Y = 1)P(Y = 1)}{P(X|Y = 1)P(Y = 1) + P(X|Y = 2)P(Y = 2)}
$$
\n
$$
= \frac{1}{1 + \frac{P(X|Y = 2)P(Y = 2)}{P(X|Y = 1)P(Y = 1)}}
$$
\n
$$
= \frac{1}{1 + \exp(\log(\frac{P(X|Y = 2)P(Y = 2)}{P(X|Y = 1)P(Y = 1)}))}
$$
\n
$$
= \frac{1}{1 + \exp(\log(P(X|Y = 2)P(Y = 2)) - \log(P(X|Y = 1)P(Y = 1)))}
$$
\n
$$
= \frac{1}{1 + \exp(-( \log(\frac{P(Y = 1)}{P(Y = 2)})) + \log(P(X|Y = 2)) - \log(P(X|Y = 1)))}
$$
\n
$$
\log(P(X|Y = 1)) = -\frac{1}{2} \ln(|\Sigma|) - \frac{D}{2} \ln(\pi) - \frac{1}{2} (x - \mu_1)^T \Sigma^{-1} (x - \mu_1)
$$
\n
$$
\log(P(X|Y = 2)) = -\frac{1}{2} \ln(|\Sigma|) - \frac{D}{2} \ln(\pi) - \frac{1}{2} (x - \mu_2)^T \Sigma^{-1} (x - \mu_2)
$$
\n
$$
\log(P(X|Y = 2)) - \log(P(X|Y = 1)) = \frac{1}{2} (x - \mu_1)^T \Sigma^{-1} (x - \mu_1) - \frac{1}{2} (x - \mu_2)^T \Sigma^{-1} (x - \mu_2)
$$

$$
\log(P(X|Y=2)) - \log(P(X|Y=1)) = (\mu_1^T - \mu_2^T)\Sigma^{-1}x + x^T\Sigma^{-1}(\mu_1 - \mu_2)
$$
  

$$
\log(P(X|Y=2)) - \log(P(X|Y=1)) = (\mu_1^T - \mu_2^T)\Sigma^{-1}x + x^T\Sigma^{-1}(\mu_1 - \mu_2)
$$
  

$$
\log(P(X|Y=2)) - \log(P(X|Y=1)) = 2(\mu_1^T - \mu_2^T)\Sigma^{-1}x + \mu_1^T\Sigma^{-1}\mu_1 - \mu_2^T\Sigma^{-1}\mu_2
$$

Plugging (5) in 4:

$$
P(Y = 1|X) = \frac{1}{1 + \exp(-(\log(\frac{P(Y=1)}{P(Y=2)}) + \mu_1^T \Sigma^{-1} \mu_1 - \mu_2^T \Sigma^{-1} \mu_2 + 2(\mu_1^T - \mu_2^T) \Sigma^{-1} x))}
$$
  
\n
$$
\P(Y = 1|X) = \frac{1}{1 + \exp(-(C) + \theta^T x))}
$$

Where

$$
C = \frac{P(Y=1)}{P(Y=2)} - \mu_1^T \Sigma^{-1} \mu_1 + \mu_2^T \Sigma^{-1} \mu_2
$$

$$
\boxed{\theta = 2(\mu_1 - \mu_2) \Sigma^{-1}}
$$

$$
(\Sigma^{-1})^T = \Sigma^{-1})
$$

since

#### Problem 2: (b2)

Given  $p(y|x)$  is logistic  $P(Y=1|X) = \frac{1}{1+\exp(-(C+\theta^T x))}$ Consider the simplification using first principles as in part $(b1)$ :

$$
P(Y = 1|X) = \frac{1}{1 + \exp(-(\log(\frac{P(Y=1)}{P(Y=2)})) + \log(P(X|Y=2)) - \log(P(X|Y=1)))}
$$

Now, consider the distribution  $P(X = x|Y = 1) = e^{-\lambda_1} \frac{\lambda_1^x}{x!}$  and  $P(X = x|Y = 2) = e^{-\lambda_2} \frac{\lambda_2^x}{x!}$ <br>  $\log(P(X|Y = 2)) - \log(P(X|Y = 1) = \lambda_1 - \lambda_2 + x(\log \frac{\lambda_1}{\lambda_2})$ and hence,

$$
P(Y = 1|X) = \frac{1}{1 + \exp(-(\log(\frac{P(Y=1)}{P(Y=2)})) + \lambda_1 - \lambda_2 + x(\log \frac{\lambda_1}{\lambda_2})))}
$$

implying it is possible to arrive at a logistic regression expression even from a poisson distribution and hence the  $p(x|y)$  need not be gaussian.

# Problem 3

$$
L(w_{i+1}, \lambda) = ||w_{i+1} - w_i||_2^2 + \lambda (w_{i+1}^T x_i) y_i
$$
  
=  $(w_{i+1} - w_i)^T (w_{i+1} - w_i) + \lambda (w_{i+1})$   

$$
\Delta_{w_{i+1}} L = 2w_{i+1} - 2w_i + \lambda x_i y_i = 0
$$
  

$$
\Delta_{\lambda} L = w_{i+1}^T x_i y_i = 0
$$
 (3.1)

3

Thus, from 3.1 and 3.2,

$$
w_{i+1} = w_i - \frac{1}{2}\lambda x_i
$$
  
where  

$$
w_{i+1}^T x_i y_i = 0
$$

# Problem 4

4

### Problem 4: (a)



#### Problem 4: (b)

From the graph above we see that pclass and 'age' might not be really informative. (Given they do not have any monotonicity). whereas the rest all variables have.

#### Problem 4: (c)





Figure 1: 4a. Montonic relationship

## Problem 4: (d)

MM: Multiple Models SM: Substituted Models IM: Individual Model



Thus, the individual model(IM) with the age column completely removed seems to have worked better than the MM. Though in training MM perfroms poorer than SM, its performance is at part with SM for the test dataset. In totality, Substituted model seems to have worked better.(considering both training and test datasets).

This is an indicative that the 'age' factor is not really informative as is also evident from part (c) above where 'age' features low in the information table.

## Problem 4: (e)

Total number of columns: 602.

# Problem 4: (f)

The method of forward selection seems to have worked well. The training accuracy increased by increasing the number of features iteratively. This also lead to an increase in test accuracy though only marginally. As evident, the training accuracy plot seems to flatten (and hence saturate) near 0.85 for around 10 features. So 10 features can be assumed to be an optimal choice for number of features.

## Problem 4: (g)





Figure 2: 4f. Training/testing accuracy v/s iteration



e parameter around  $0.1$ and gives an accuracy of 0.94. (Also very low values of the slope parameter seems to hit a local minima) It does seem to be converging in a stable way. glmfit accuracy: 0.984733

#### Problem 4: (h)

Number of iterations: 25 Accuracy: 0.583969 glmfit accuracy: 0.984733 Newton's methods implementation seems to be buggy somewhere.