

MATH-650 Assignment 6

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Problem 19

```
pH.data <- read.csv('case0702.csv', header=T)
logT <- log(pH.data$Time)
pH.data$logT <- logT
n <- nrow(pH.data)
```

Simple linear regression model for $\log(pH)$ is given by:

$$pH = \beta_0 + \beta_1 \log(T)$$

Part (a)

```
fit <- lm(pH~logT, data=pH.data)
s <- summary(fit)
b0 <- fit$coefficients[1]
b1 <- fit$coefficients[2]
se0 <- s$coefficients[3]
se1 <- s$coefficients[4]
t0 <- s$coefficients[5]
t1 <- s$coefficients[6]
p0 <- s$coefficients[7]
p1 <- s$coefficients[8]
```

Thus, $\beta_0 = 6.983626$ ($S.E = 0.048532$, $p - value = 6.0839895 \times 10^{-15}$) and $\beta_1 = -0.7256578$ ($S.E = 0.0344263$, $p - value = 2.6951582 \times 10^{-8}$)

Part (b)

```
Xbar <- mean(log(pH.data$Time))#1.190
sx2 <- var(log(pH.data$Time))#0.6344
mu <- b0 +b1*log(5)
```

Thus, $\hat{\mu}\{Y|\log(5)\} = 5.8157249$

Part (c)

```
sigmahat <- 0.08226
se = sigmahat *sqrt(1/n+(log(5)-Xbar)^2/(n-1*sx2))
```

$SE[\hat{\mu}\{Y|\log(5)\}] = 0.0283496$

Problem 25

Part (a)

$$\begin{aligned}SS(\beta_0, \beta_1) &= \sum_{i=1}^N (Y_i - \beta_0 - \beta_1 X_i)^2 \\ \frac{\partial SS}{\partial \beta_0} &= -2 \sum_{i=1}^N (Y_i - \beta_0 - \beta_1 X_i) \\ \frac{\partial SS}{\partial \beta_1} &= -2 \sum_{i=1}^N (Y_i - \beta_0 - \beta_1 X_i) X_i \\ \frac{\partial SS}{\partial \beta_0} &= -2 \sum_{i=1}^N (Y_i - \beta_0 - \beta_1 X_i) = 0 \\ \sum_{i=1}^N Y_i &= \sum_{i=1}^N (\beta_0 + \beta_1 X_i) \\ \sum_{i=1}^N Y_i &= \beta_0 n + \beta_1 \sum_{i=1}^N X_i \\ \frac{\partial SS}{\partial \beta_1} &= -2 \sum_{i=1}^N (Y_i - \beta_0 - \beta_1 X_i) X_i = 0 \\ \sum_{i=1}^N Y_i X_i &= \sum_{i=1}^N (\beta_0 + \beta_1 X_i) X_i \\ \sum_{i=1}^N Y_i X_i &= \beta_0 \sum_{i=1}^N X_i + \sum_{i=1}^N \beta_1 X_i^2\end{aligned}$$

Part (b)

$$\hat{\beta}_0 = \bar{Y} - \hat{\beta}_1 \bar{X}$$

And,

$$\begin{aligned}
\hat{\beta}_1 &= \frac{\sum_{i=1}^N (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^N (X_i - \bar{X})^2} \\
&= \frac{\sum_{i=1}^N (X_i Y_i - \bar{X} Y_i + \bar{X} \bar{Y} - X_i \bar{Y})}{\sum_{i=1}^N (X_i^2 - 2\bar{X} X_i + \bar{X}^2)} \\
&= \frac{\sum_{i=1}^N X_i Y_i - n\bar{X}\bar{Y}}{\sum_{i=1}^N X_i^2 - n\bar{X}^2}
\end{aligned}$$

Thus,

$$\begin{aligned}
\hat{\beta}_1 \sum_{i=1}^N X_i^2 - n\hat{\beta}_1 \bar{X}^2 &= \sum_{i=1}^N X_i Y_i - n\bar{X}\bar{Y} \\
\hat{\beta}_1 \sum_{i=1}^N X_i^2 + n\bar{X}\bar{Y} - n\hat{\beta}_1 \bar{X}^2 &= \sum_{i=1}^N X_i Y_i \\
\hat{\beta}_1 \sum_{i=1}^N X_i^2 + n\bar{Y}(\bar{X} - \hat{\beta}_1 \bar{X}) &= \sum_{i=1}^N X_i Y_i
\end{aligned}$$

Substituting (3) in (4), we get

$$\begin{aligned}
\hat{\beta}_1 \sum_{i=1}^N X_i^2 + n\bar{X}(\bar{Y} - \hat{\beta}_1 \bar{X}) &= \sum_{i=1}^N X_i Y_i \\
\hat{\beta}_1 \sum_{i=1}^N X_i^2 + n\bar{X}(\hat{\beta}_0) \bar{X} &= \sum_{i=1}^N X_i Y_i \\
\hat{\beta}_1 \sum_{i=1}^N X_i^2 + \hat{\beta}_0 \sum_{i=1}^N X_i &= \sum_{i=1}^N X_i Y_i
\end{aligned}$$

Thus, (5) is same as (2)

Now, From (3)

$$\begin{aligned}
\hat{\beta}_0 &= \bar{Y} - \hat{\beta}_1 \bar{X} \\
\bar{Y} &= \hat{\beta}_0 + \hat{\beta}_1 \bar{X} \\
\frac{\sum_{i=1}^N Y_i}{N} &= \hat{\beta}_0 + \hat{\beta}_1 \frac{\sum_{i=1}^N X_i}{N} \\
\sum_{i=1}^N Y_i &= n\hat{\beta}_0 + \hat{\beta}_1 \sum_{i=1}^N X_i
\end{aligned}$$

Thus, (6) is same as (1)

Consider second order differentials:

$$\begin{aligned}
\frac{\partial^2 SS}{\partial \beta_0^2} &= -2 * -n = 2n \\
\frac{\partial^2 SS}{\partial \beta_0 \partial \beta_1} &= 2 \sum_{i=1}^N X_i \\
\frac{\partial^2 SS}{\partial \beta_1 \partial \beta_0} &= 2 \sum_{i=1}^N X_i \\
\frac{\partial^2 SS}{\partial \beta_1^2} &= 2 \sum_{i=1}^N X_i^2 \\
\text{Hessian}(H) &= \begin{bmatrix} 2 & 2 \sum_{i=1}^N X_i \\ 2 \sum_{i=1}^N X_i & 2 \sum_{i=1}^N X_i^2 \end{bmatrix} \\
&2n > 0 \text{ and,} \\
\det(H) &= 4(n \sum_{i=1}^N X_i^2 - (\sum_{i=1}^N X_i)^2) = 4n(n-1)\text{Var}(X) > 0 \text{ always}
\end{aligned}$$

Thus, the values of $\hat{\beta}_0, \hat{\beta}_1$ indeed guarantee a minima since the Hessian is positive definite.

Part (c)

Problem 13

Intercept = 0.3991 Standard Error = 0.1185 df = 22 $t_{0.975,22} = 2.073$

Upper limit = 0.3991 + 2.073 * 0.1185 = 0.6447505 Lower limit = 0.3991 - 2.073 * 0.1185 = 0.1534495

Thus, the 95% CI for Intercept is [0.1534495, 0.6447505]

Also following is the R code:

```

intercept <- 0.3991
se <- 0.1185
df <- 22
t975 <- qt(0.975,df)
limit.upper <- intercept + t975*se
limit.lower <- intercept - t975*se
limit.upper
## [1] 0.644854

```

```
limit.lower
```

```
## [1] 0.153346
```