

Lecture 4: Sep 1, 2016

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Example: **Branching Processes** Day 0: ... (organisms)

Day 1: Y_1, Y_2 , (Number of offsprings of i^{th} bacteria)

$\{Y_i\}$ are iid.

$Y_i \in \{0, 2, 4\}$ $p_0 + p_2 + p_4 = 1$

Day n \implies M bacteria ($X_n = m$)

$X_n = \#$ of bacteria on day n

$P(X_{n+1} = j | X_n = M) = P(Y_1 + Y_2 + \dots + Y_M = j | X_n = M) = P(X_{n+1} = j | X_n = M, X_{n-1} = a, \dots, X_0 = 2)$

so a branching process is markov chain

Multistep Transition Probability

$P(X_{n+m} = j | X_n = 1, X_{n-1}) = P(X_m = j | X_0 = i)$ m - step transition probabilities does on depend on past given the current state.

Gambler's ruin: time homegenous $P(X_1 = 1 | X_0 = 0) = P(X_{101} = 1 | X_{100} = 0)$

Example: $S = \{1, 2, 3\}$

$$p = \begin{pmatrix} 0.8 & 0.1 & 0.1 \\ 0.1 & 0.8 & 0.1 \\ 0.1 & 0.1 & 0.8 \end{pmatrix}$$

$P(X_2 = 1, X_1 = 3 | X_0 = 2) = P(X_2 + 1 | X_1 = 3)P(X_1 = 3 | X_0 = 2)$

$P(A \cap B) = P(A|B)P(B) \implies P(A \cap B|C) = P(A|B \cap C)P(B|C)$

$P(X_2 = 1, X_1 = 3 | X_0 = 2) = P(X_2 = 1 | X_1 = 3)P(X_1 = 3 | X_0 = 2) = 0.1 * 0.1 = 0.01$

$P(X_2 = 1 | X_0 = 2) = \sum_i P(X_2 = 1, X_1 = i | X_0 = 2)P(X_i = i | X_0 = 2)$

$P(X_2 = j | X_0 = i)$ denoted by $p^2(i, j)$

$p^2(i, j) = \sum_{k \in S} P(X_2 = j, X_1 = k | X_0 = i) = \sum_{k \in S} P(X_2 = j | X_1 = k)P(X_1 = k | X_0 = i) = \sum_{k \in S} p(k, j)p(i, k)$

$p^2 = p.p$ Two step is product of one step.

So m step transition probabilities are denoted by $p^m(i, j) = P(X_m = j | X_0 = i)$

m - step transition probability matrix is the $m^{[th]}$ power of one step transition probability matrix.

4.0.1 Champan Kplmogorov equations

$$p^{m+n}(i, j) = \sum_{k \in S} p^m(i, k) p^n(k, j)$$

$$P(X_{m+n} = j | X_0 = i)$$

Proof:

$$\begin{aligned} P(X_{n+m} = k | X_0 = i) &= \sum_{k \in S} P(X_{n+m} = j, X_m = k | X_0 = i) \\ &= \sum_{k \in S} P(X_{n+m} = j | X_m = k, X_0 = i) P(X_m = k | X_0 = i) \\ &= \sum_{k \in S} P(X_{n+m} = j | X_m = k) P(X_m = k | X_0 = i) \\ &= \sum_{k \in S} p^n(k, j) p^m(i, k) \end{aligned}$$

Notation

For some event A , $P(A | X_0 = x) = P_x(A)$

For some RV Y , $E[Y | X_0 = x] = E_x[Y]$

Stopping Time

Definition Consider a RV T , takes a values in the set $\{0, 1, 2, 3\}$. We say that T is a stopping time with respect to the stochastic process. $\{X_n, n \geq 0\}$ if the value of I_N [INCOMPLETE]

$\{X_n, n \geq 0\}$ – Gambler's ruin problem

$$T = \min\{k \geq 0 : X_k \geq 5\}$$

Example: $X_0 = 10$ then $T = 0$

Calculate $I_{\{T_n\}}$ from X_0, X_1, \dots, X_n

$$\text{If } \{T = 0\} \implies \{X_0 \geq 5\}$$

$$\{T = 1\} \implies \{X_0 < 5, X_1 \geq 5\}$$

$$\{T = n\} \implies \{X_0 < 5, X_1 < 5, \dots, X_{n-1} < 5, X_n \geq 5\}$$

T is a stopping time

Strong markov property:

Suppose T is a stopping time with respect to the Markov Chain, $\{X_n, n \geq 0\}$ Given that stopping time $T = n$ and $X_T = y$

Any other information about X_0, X_1, \dots, X_{T-1} is irrelevant for predicting future.

$$P(X_{T+1} = j | T = n, X_T = y) = P(X_1 = j | X_0 = y) = p(y, j)$$

$$P(X_{T+1} = j | T = n, X_T = y, X_{T-1} = x, \dots, X_0 = z) = P(X_{T+1} = j | T = n, X_T = y) P(y, j)$$

$$P(X_{T+m} = j | T = n, X_T = y, X_{T-1} = x) = P(X_{T+m} = j | T = n, X_T = y) = P(X_m = j | X_0 = y) = p^m(y, j)$$