

MATH 542 Homework 9

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Problem 1

Since \mathbf{X} is full rank, the least square estimate of β for the linear model $\mathbf{Y} = \mathbf{y}\beta + \epsilon$ then $\hat{\beta} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y}$ also this is an unbiased estimator of β that is $E[\hat{\beta}] = \beta$

$$\begin{aligned} a'\beta &= a'E[\hat{\beta}] \\ &= a'E[(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y}] \\ &= E[a'(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y}] \\ &= E[b'\mathbf{y}] \end{aligned}$$

Hence when \mathbf{X} is full rank, we have can take $b' = a'(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'$ making every $a'\beta$ estimable

For each individual β_i to be estimable, we set $a' = (0 \ 0 \ \dots \ 0 \ 1 \ 0 \dots 0)$ setting $a_i = 1$ remaining zero.

Problem 2

Since each $a'_i\beta$ is estimable: $a'_i\beta = E[b'_iY]$ Consider $\lambda_i \in \mathcal{R}$

$$\begin{aligned} \lambda_1 a'_1\beta + \lambda_2 a'_2\beta + \dots + \lambda_n a'_n\beta &= \lambda_1 E[b'_1Y] + \lambda_2 E[b'_2Y] + \dots + \lambda_n E[b'_nY] \\ &= \sum_i E[\lambda_i b'_i Y] \\ &= \sum E[c'Y] \text{ where } c' = \lambda \mathbf{b}' \end{aligned}$$

where $\lambda = (\lambda_1 \ \lambda_2 \ \dots \ \lambda_n)$ and $b' = (b_1 \ b_2 \ \dots \ b_n)$

Hence linear combination of $a'_i\beta$ is also estimable.

Problem 3

$$\begin{aligned} X &= \begin{pmatrix} 1 & 1 & 3 \\ 1 & 1 & -2 \\ 1 & 1 & 0 \\ 1 & 0 & 0 \end{pmatrix} \beta &= \begin{pmatrix} \beta_1 \\ \beta_2 \\ \beta_3 \end{pmatrix} \\ \beta_1 &= (1 \ 1 \ 0) \beta \\ \beta_1 - \beta_2 &= (1 \ -1 \ 0) \beta \\ 5\beta_1 + 3\beta_2 + 9\beta_3 &= (5 \ 3 \ 9) \beta \end{aligned}$$

X has full column rank and hence we can make use of the theorem we proved in Problem 1 to say that all three cases (a, b, c) are indeed estimable, that is since X is full rank, every $a'\beta$ is estimable (One particular case of the corollary proved in Problem 1 here is the case a where $a' = (1 \ 0 \ 0)$)

Problem 4

Problem 4.a

Define $\mathbf{y} = (Y_1 \ Y_2 \ Y_3)'$, $\tau = (\tau_1 \ \tau_2 \ \tau_3)$, $\epsilon = (\epsilon_1 \ \epsilon_2 \ \epsilon_3)$

$$\text{Thus, } \mathbf{y} = \begin{pmatrix} 1 & 1 & 1 \\ 1 & 0 & 1 \\ 0 & 1 & 0 \end{pmatrix} \tau + \epsilon$$

Rank of design matrix is 2

Problem 4.b

$E[b'Y] = a'\beta = b'X\beta$ iff $a' = b'X$ or $a = Xb'$ τ_2 is estimable:

$$\begin{aligned} \tau_2 &= a'\beta = (0 \ 1 \ 0) \tau \\ &= EY_3 \\ &= E[(0 \ 0 \ 1)] \end{aligned}$$

$\tau_1 = (1 \ 0 \ 0) \tau$ so we need to find b such that $(1 \ 0 \ 0)' = X'b$ It is clear to see no b does not exist.

$\tau_2 = (0 \ 1 \ 0) \tau$ is estimable because $b = (0 \ 1 \ 0)$

τ_3 is not estimable because of symmetry with τ_1

Problem 4.c

$$\tau_1 - 2\tau_2 + \tau_3 = (1 \ -2 \ 1) \tau = \begin{pmatrix} 1 & 1 & 0 \\ 1 & 0 & 1 \\ 0 & 1 & 0 \end{pmatrix} \begin{pmatrix} 0 \\ 1 \\ -2 \end{pmatrix} \text{ and hence it is estimable}$$

Problem 4.d

A possible unbiased estimator of $\tau_1 - 2\tau_2 + \tau_3 = E[Y_2 - 2Y_3]$ i.e $Y_2 - 2Y_3$ which is not necessarily BLUE.

For BLUE, we simply take the OLS estimate of β as $\hat{\beta} = \mathbf{X}'\mathbf{X}^{-1}\mathbf{X}'\mathbf{y}$ and from Gauss-Markov model BLUE follows.

$$\text{Generalized inverse (using } R) \ \mathbf{X}'\mathbf{X}^{-1} = \begin{pmatrix} \frac{1}{6} & \frac{-1}{6} & \frac{1}{6} \\ \frac{-1}{6} & \frac{2}{3} & \frac{-1}{6} \\ \frac{1}{6} & \frac{-1}{6} & \frac{1}{6} \end{pmatrix}$$

$$\text{so } \hat{\beta} = \mathbf{X}'\mathbf{X}^{-1}\mathbf{X}'\mathbf{y} = \begin{pmatrix} \frac{1}{6} & \frac{1}{3} & \frac{-1}{6} \\ \frac{1}{3} & \frac{-1}{3} & \frac{2}{3} \\ \frac{1}{6} & \frac{1}{3} & \frac{-1}{6} \end{pmatrix} \mathbf{y}$$

$$\text{and hence } \tau_1 - 2\tau_2 + \tau_3 = (1 \ -2 \ 1) \begin{pmatrix} \frac{1}{6} & \frac{1}{3} & \frac{-1}{6} \\ \frac{1}{3} & \frac{-1}{3} & \frac{2}{3} \\ \frac{1}{6} & \frac{1}{3} & \frac{-1}{6} \end{pmatrix} \mathbf{y} = \frac{-Y_1 + 4Y_2 - 5Y_3}{3}$$

$$\text{Thus BLUE of } \tau_1 - 2\tau_2 + \tau_3 : \frac{-Y_1 + 4Y_2 - 5Y_3}{3}$$