

# **MATH-605: Homework # 4**

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## 5.4.12

$$\begin{aligned}
\mathbb{E} \exp \lambda \epsilon A &= \frac{1}{2} (\exp \lambda A + \exp -\lambda A) \\
\exp A &= I + A + \frac{A^2}{2!} + \frac{A^3}{3!} + \dots \\
\exp -A &= I - A + \frac{A^2}{2!} - \frac{A^3}{3!} + \dots \\
\frac{1}{2} (\exp \lambda A + \exp -\lambda A) &= I + \frac{\lambda^2 A^2}{2} + \frac{\lambda^4 A^4}{4!} + \frac{\lambda^6 A^6}{6!} + \dots \\
&= I + \frac{\lambda^2 A^2}{2} + \frac{\lambda^4 A^4}{4 * 3} + \frac{\lambda^6 A^6}{8 * 90} + \dots \\
&\leq 1 + (\lambda^2 A^2 / 2) + \frac{((\lambda^2 A^2) / 2)^2}{2} + \frac{(\lambda^2 A^2 / 2)^3}{3!} + \dots \\
&= \exp \lambda^2 A^2 / 2
\end{aligned}$$

Then define  $X = \sum_{i=1}^N \epsilon_i A_i$ , then following from 5.14 :

$$\begin{aligned}
\mathbb{P}\{\lambda_{max}(S) \geq t\} &= \mathbb{P}\{e^{\lambda \lambda_{max}(S)} \geq t\} \\
&\leq e^{-\lambda t} \mathbb{E} e^{\lambda \lambda_{max}(S)}
\end{aligned}$$

Define  $E = \mathbb{E} \lambda_{max}(e^{\lambda S})$ . By the bound on maximum eigen value of  $e^{\lambda S}$ :  $E \leq \mathbb{E} \text{tr} e^{\lambda S}$

Applying Lieb's inequality:

$$E \leq \mathbb{E} \text{tr} e^{\lambda S}$$

$$E \leq \mathbb{E} \text{tr} \exp \left[ \sum_{i=1}^{N-1} \lambda X_i + \lambda X_N \right]$$

$$E \leq \mathbb{E} \text{tr} \exp \left[ \sum_{i=1}^{N-1} \lambda X_i + \log \mathbb{E} e^{\lambda X_N} \right] \quad \text{Conditioning on } X_{i=1}^{N-1} \text{ and using Lemma 5.4.9}$$

$$E \leq \text{tr} \exp \left[ \sum_{i=1}^N \log \mathbb{E} e^{\lambda X_i} \right] \quad \text{Repeating above step and using Lemma 5.4.9 } N \text{ times}$$

$$\leq \text{tr} \exp \left[ \sum_{i=1}^N \frac{\lambda_i^2 A_i^2}{2} \right] \quad \because \mathbb{E} \exp \lambda \epsilon A \leq \exp \lambda^2 A^2 / 2$$

$$\leq n \cdot \lambda_{max} (\exp \sum_{i=1}^N \lambda_i^2 A_i^2 / 2) \quad \because \text{trace is sum of } n \text{ eigen values}$$

$$\leq n \exp \left\{ n \lambda_{max}^2 / 2 \sum_{i=1}^n A_i^2 \right\} \quad \because \lambda_i \leq \lambda_{max} \forall i \in [1, N]$$

$$= n \exp \{ \lambda_{max}^2 \sigma^2 / 2 \}$$

$$\mathbb{P}\{\lambda_{max}(S) \geq t\} \leq n \cdot \exp \{ -\lambda t + \lambda_{max}^2 \sigma^2 / 2 \} \quad \text{substituting for } E$$

Differentiating  $\exp \{ -\lambda t + \lambda_{max}^2 \sigma^2 / 2 \}$  wrt  $\lambda$  gives:  $\exp \{ -\lambda t + \lambda_{max}^2 \sigma^2 / 2 \} * (-t + \lambda_{max} \sigma^2 / 2)$  Thus  $\lambda = \frac{t}{\sigma^2}$  and hence:

$$E \leq n \cdot \exp \left\{ -\frac{t^2}{2\sigma^2} \right\}$$

### 5.4.15

Consider dilation of  $X$  as  $Y := \begin{pmatrix} 0 & X_i^T \\ X_i & 0 \end{pmatrix}$  Then  $Y^2 = \begin{pmatrix} X_i^T X_i & 0 \\ 0 & X_i X_i^T \end{pmatrix}$  Then  $\sigma^2 = \|\sum_{i=1}^N Y_i^2\| = \left\| \begin{pmatrix} \sum_{i=1}^N X_i^T X_i & 0 \\ 0 & \sum_{i=1}^N X_i X_i^T \end{pmatrix} \right\| = \max\{\|\sum_{i=1}^N X_i^T X_i\|, \|\sum_{i=1}^N X_i X_i^T\|\}$  Applying Matrix Bernstein's inequality from theorem 5.4.1 to the dilation  $Y$  of  $X$  give:

$$\begin{aligned} P\{\|\sum_{i=1}^N X_i\| \geq t\} &= P\{\|\sum_{i=1}^N Y_i\| \geq t\} \\ &\leq 2(m+n) \exp -\frac{t^2/2}{\sigma^2 + Kt/3} \end{aligned}$$

where  $\sigma^2 = \max\{\|\sum_{i=1}^N X_i^T X_i\|, \|\sum_{i=1}^N X_i X_i^T\|\}$

### 5.6.6

Frame  $u_i$  obeys's approximate Parseva;s identity:  $\exists A, B > 0$  such that

$$A\|x\|_2^2 \leq \sum_{i=1}^N \langle u_i, x \rangle \leq B\|x\|_2^2 \forall x \in \mathbb{R}^n$$

$u_i$  is tight when  $A = B$ . Also, from problem 3.3.9 we have  $\{u_i\}_{i=1}^N$  is tight when  $\sum_{i=1}^n u_i u_i^T = AI_n$  Consider random sample  $\{v_i\}_{i=1}^m$  of  $\{u_i\}$  from remark 5.6.2 we see that

$$\begin{aligned} E\|\sum_{i=1}^m v_i v_i^T - \sum_{i=1}^N u_i u_i^T\| &\leq \epsilon \|\sum_{i=1}^N u_i u_i^T\| \\ &= \epsilon \|A\| \end{aligned}$$

Hence  $\{v_i\}_{i=1}^m$  has a good frame bound.

### 6.1.6

$EF(\sum_{i \neq j} a_{ij} f(X_i, X_j)) \leq E(4 \sum_{i \neq j} a_{ij} f(X_i, X_j'))$ . For this to hold,  $f$  should be measurable.

If this holds then theorem 6.1.1 is implied by taking  $f(X_i, X_j) = X_i X_j$  for matrix  $X_i$  and theorem 6.1.4 is implied by considering  $f(X_i, X_j) = X_i X_j^T$  for vectors  $X_i, X_j$ .

**6.3.4**

To prove:  $\mathbb{E} \left\| \sum_{i=1}^N X_i - \sum_{i=1}^N \mathbb{E} X_i \right\| \leq 2 \mathbb{E} \left\| \sum_{i=1}^N \epsilon_i X_i \right\|$

$$\begin{aligned} \mathbb{E} \left\| \sum_{i=1}^N X_i - \sum_{i=1}^N \mathbb{E} X_i \right\| &\leq \mathbb{E} \left\| \sum_{i=1}^N X_i \right\| + \left\| \sum_{i=1}^N \mathbb{E} X_i \right\| && \text{triangular inequality} \\ &\leq \mathbb{E} \left\| \sum_{i=1}^N \epsilon_i X_i \right\| + \left\| \sum_{i=1}^N \mathbb{E} \epsilon_i X_i \right\| && \text{same distribution} \\ &\leq \mathbb{E} \left\| \sum_{i=1}^N \epsilon_i X_i \right\| + \mathbb{E} \left\| \sum_{i=1}^N \epsilon_i X_i \right\| && \text{Jensen's inequality} \\ &= 2 \mathbb{E} \left\| \sum_{i=1}^N \epsilon_i X_i \right\| \end{aligned}$$

## 6.3.5

Consider  $Y_i$  to be independent copies of  $X_i$

$$\begin{aligned} \left\| \sum_{i=1}^N X_i \right\|_F &= \left\| \sum_{i=1}^N X_i - \sum_{i=1}^N \mathbb{E}Y_i \right\|_F && \because \mathbb{E}Y_i = 0 \\ &\leq \mathbb{E}_Y \left\| \sum_{i=1}^N (X_i - Y_i) \right\|_F \end{aligned}$$

$$\begin{aligned} F\left(\left\| \sum_{i=1}^N X_i \right\|_F\right) &\leq F\left(\mathbb{E}_Y \left\| \sum_{i=1}^N (X_i - Y_i) \right\|_F\right) \\ &\leq \mathbb{E}_Y F\left(\left\| \sum_{i=1}^N (X_i - Y_i) \right\|_F\right) && \text{Jensen's} \end{aligned}$$

$$\begin{aligned} \mathbb{E}_{X,Y} F\left(\left\| \sum_{i=1}^N X_i \right\|_F\right) &\leq \mathbb{E}_X \mathbb{E}_Y F\left(\left\| \sum_{i=1}^N (X_i - Y_i) \right\|_F\right) \\ &\leq \mathbb{E}_{X,Y} F\left(\left\| \sum_{i=1}^N (X_i - Y_i) \right\|_F\right) && \text{Using Fubini's} \end{aligned}$$

Now, for rademacher  $\epsilon_i$

$$\begin{aligned} \mathbb{E}_{X,Y} F\left(\left\| \sum_{i=1}^N (X_i - Y_i) \right\|_F\right) &\leq \mathbb{E}_\epsilon \mathbb{E}_{X,Y} F\left(\left\| \sum_{i=1}^N \epsilon_i (X_i - Y_i) \right\|_F\right) \\ &\leq \mathbb{E}_{X,Y} F\left(\left\| \sum_{i=1}^N \epsilon_i X_i \right\|_F\right) + \mathbb{E}_{X,Y} F\left(\left\| \sum_{i=1}^N \epsilon_i Y_i \right\|_F\right) \\ &\leq 2 \mathbb{E}_{X,Y} F\left(\left\| \sum_{i=1}^N \epsilon_i X_i \right\|_F\right) \end{aligned}$$

which is the upper bound.

For lower bound:

$$\begin{aligned} \mathbb{E} F\left(\frac{1}{2} \left\| \sum_{i=1}^N \epsilon_i X_i \right\|_F\right) &= \mathbb{E}_\epsilon \mathbb{E} F\left(\frac{1}{2} \left\| \sum_{i=1}^N \epsilon_i X_i - \sum_{i=1}^N \mathbb{E}Y_i \right\|_F\right) \\ &\leq \mathbb{E} \left\| \frac{1}{2} \sum_{i=1}^N \epsilon_i (X_i - Y_i) \right\|_F \\ &\leq \mathbb{E} \left\| \frac{1}{2} \sum_{i=1}^N \epsilon_i X_i \right\|_F + \mathbb{E} \left\| \frac{1}{2} \sum_{i=1}^N \epsilon_i Y_i \right\|_F \\ &= \mathbb{E} \left\| \frac{1}{2} \sum_{i=1}^N X_i \right\|_F + \mathbb{E} \left\| \frac{1}{2} \sum_{i=1}^N Y_i \right\|_F && \text{Same distribution} \\ &\leq \frac{1}{2} \left( \mathbb{E} \left\| \sum_{i=1}^N X_i \right\|_F + \mathbb{E} \left\| \sum_{i=1}^N Y_i \right\|_F \right) \\ &= \mathbb{E} \left\| \sum_{i=1}^N X_i \right\|_F \end{aligned}$$

which is the LHS of the whole inequality.

## 6.5.4

$\hat{X}$  is best approximation to  $p^{-1}Y$  hence  $\|\hat{X} - p^{-1}Y\| \leq \|p^{-1}Y - X\|$

$$\|\hat{X} - X\| \leq \|\hat{X} - p^{-1}Y\| + \|p^{-1}Y - X\| \quad \text{Trinagular inequality}$$

$$\|\hat{X} - X\| \leq 2\|p^{-1}Y - X\| \quad \because \text{Assumption above}$$

$$= \frac{2}{p}\|Y - pX\|$$

$$(Y - pX)_{ij} = (\delta_{ij} - p)X_{ij} + \delta_{ij}v_{ij}$$

Consider,

$$\begin{aligned} \|(Y - pX)_i\|_2^2 &= \sum_{j=1}^n (\delta_{ij} - p)X_{ij} + \delta_{ij}v_{ij})^2 \\ &\leq \sum_{j=1}^n ((\delta_{ij} - p)\|X\|_\infty + \delta_{ij}\|v\|_\infty)^2 \end{aligned}$$

$$\mathbb{E} \max_{j=1}^n (\delta_{ij} - p)^2 \leq Cpn \text{ Using 2.8.3}$$

$$\mathbb{E} \max_{j=1}^n (\delta_{ij})^2 \leq Cpn$$

$$\begin{aligned} \implies \mathbb{E} \max \|(Y - pX)_i\|_2^2 &\leq \mathbb{E} \sum_{j=1}^n ((\delta_{ij} - p)\|X\|_\infty + \delta_{ij}\|v\|_\infty)^2 \\ &= \mathbb{E} \sum_{j=1}^n ((\delta_{ij} - p)^2\|X\|_\infty^2 + \delta_{ij}^2\|v\|_\infty^2 + (\delta_{ij} - p)\delta_{ij}\|X\|_\infty\|v\|_\infty)^2 \end{aligned}$$

$$\implies \mathbb{E} \max \|(Y - pX)_i\|_2 \leq \sqrt{pn}\|X\|_\infty + \sqrt{pn}\|v\|_\infty$$

Using 6.4.2 .

$$\mathbb{E}\|(Y - pX)\| \leq C\sqrt{\log n}(\mathbb{E} \max \|(Y - pX)_i\| + \mathbb{E} \max \|(Y - pX)^j\|)$$

Thus,

$$\mathbb{E}\|(\hat{X} - X)\| \leq \sqrt{\frac{n \log n}{p}}\|X\|_\infty + \sqrt{\frac{n \log n}{p}}\|v\|_\infty$$

**6.6.5**

$$\begin{aligned}\mathbb{E}\|g\|_\infty &= \mathbb{E}\{\max_{i \leq n} g_i\} \\ e^{s\mathbb{E}[\max_{i \leq n} g_i]} &\leq \mathbb{E}[e^{s \max_{i \leq n} g_i}] && \text{Jensen's inequality} \\ &= \mathbb{E}[\max_{i \leq n} e^{s g_i}] \\ &\leq \sum_{i=1}^n \mathbb{E}[e^{s g_i}] \\ &\leq n e^{\sigma^2 s^2 / 2} && \text{Using mgf of } g \\ \Rightarrow \mathbb{E}\{\max_{i \leq n} g_i\} &\leq \frac{\ln}{s} + \frac{s\sigma^2}{2}\end{aligned}$$

Differentiating  $\frac{\ln}{s} + \frac{s\sigma^2}{2}$  wrt  $s$  gives  $s = \sqrt{\frac{2 \ln n}{\sigma^2}}$  and hence

$$\begin{aligned}\mathbb{E}\|g\|_\infty &= \mathbb{E}\{\max_{i \leq n} g_i\} \\ &\leq \sqrt{2}\sigma\sqrt{\ln n}\end{aligned}$$